

On the average minimum size of a manipulating coalition

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Abstract: We study the asymptotic average minimum manipulating coalition size as a characteristic of quality of a voting rule and show its serious drawback. We suggest using the asymptotic average threshold coalition size instead. We prove that, in large electorates, the asymptotic average threshold coalition size is maximised among all scoring rules by the Borda rule when the number m of alternatives is 3 or 4, and by $\lfloor m/2 \rfloor$ -approval voting when $m \geq 5$.

Key words: Borda rule, approval voting, coalition, manipulability

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1 Introduction

In 1973–75 Gibbard and Satterthwaite published a fundamental impossibility theorem which states that every non-dictatorial social choice function, whose range contains at least three alternatives, at certain profiles can be manipulated by a single individual [9, 19]. After that, the natural question arose: if there are no perfect rules, which ones are the best, i.e. least manipulable? The answer to this question cannot be given in absolute terms. It stipulates introducing a certain index of manipulability of the rule and a certain model for the population. The answer may depend both on the choice of the index and on the choice of the model.

To answer this question, various indices of manipulability of voting rules have been introduced and studied [1, 5, 11, 17, 18] and various models introduced [2, 18]. Among the models the following two have gained the lion's share of attention. The first one, IC (Impartial Culture) conjecture, assumes that all voters are independent. The second one, the IAC (Impartial Anonymous Culture) conjecture, assumes some degree of dependency.

Among the indices, the probability of obtaining a manipulable profile either for an individual or for a coalition has attracted most attention, both for the IC and for the IAC conjectures [5, 14, 17].

In some cases, these indices have been calculated exactly, mainly in the three-alternative case [8, 14, 15, 16]. In others, they have been estimated in computer simulations [1, 5, 11, 17, 21].

The probability of manipulation has been especially well-studied for the important class of scoring voting rules, and significant progress has been made in comparing them. In his seminal paper [18], Saari showed that in his “geometric” model, Borda is the least manipulable for the three-alternative case in relation to micro manipulation, but this does not extend to the case of four alternatives. Kim and Roush [12], on the other hand, proved that, asymptotically, when the number of voters tends to infinity Borda becomes coalitionally manipulable with probability 1.

Some other characteristics have also been used, mostly in computer simulations [1, 5, 17]. In this paper, we originally aimed to compare scoring rules under the IC, using the average minimum size of the coalition capable of manipulation as the principal characteristic of manipulability of the rule. The greater this characteristic, the better the rule. This characteristic was first introduced by Chamberlin [5], where he estimated it for four different rules by means of computer simulation.

To our surprise we discovered that while Chamberlin's characteristic is meaningful for most classical rules, it is biased towards antiplurality, and the closer the rule gets to the antiplurality, the larger is its average minimum

coalition size. Hence, there are no optimal rules in relation to this characteristic at all. We suggest using the asymptotic average threshold coalition size instead. This new characteristic coincides with the average minimum coalition size on classical rules but does not have the bias towards antiplurality. We show how to calculate the limiting value of this characteristic when $n \rightarrow \infty$, and discuss which rule is optimal for large electorates. We prove that, among all scoring rules, Borda is the optimal rule in this sense in three- and four-alternative elections. When the number of alternatives m is five or more, $\lfloor m/2 \rfloor$ -approval voting is optimal, where for each voter most preferred $\lfloor m/2 \rfloor$ alternatives are considered approved.

2 Scoring Rules

Let A and N be two finite sets of cardinality m and n respectively. The elements of $A = \{a_1, \dots, a_m\}$ will be called alternatives, the elements of N agents. We will denote agents as $1, 2, \dots, n$. We assume that the agents have preferences over the set of alternatives. By $\mathcal{L} = \mathcal{L}(A)$ we denote the set of all linear orders on A ; they represent the preferences of agents over A . The elements of the Cartesian product

$$\mathcal{L}(A)^n = \mathcal{L}(A) \times \dots \times \mathcal{L}(A) \quad (n \text{ times})$$

are called profiles. They represent the collection of preferences of the society N . If a linear order $R_i \in \mathcal{L}(A)$ represents the preferences of the i -th agent, then by aR_ib , where $a, b \in A$, we denote that this agent prefers a to b . A family of mappings $F_n: \mathcal{L}(A)^n \rightarrow A$, $n \in \mathbb{N}$, is called a social choice function (SCF).

Definition 1. Let $\vec{R} = (R_1, \dots, R_n)$ be a profile. We say that a profile \vec{R}' occurred as a result of change of mind of agents of a coalition $C = \{i_1, \dots, i_k\} \subseteq N$, if agents of C who previously submitted linear orders R_{i_1}, \dots, R_{i_k} now submit linear orders $R'_{i_1}, \dots, R'_{i_k}$ while the remaining voters submit their original linear orders.

Definition 2. Let F be an SCF and $C = \{i_1, \dots, i_k\} \subseteq N$ be a coalition. We say that a profile \vec{R} is manipulable for F by the coalition C , if there is a profile \vec{R}' , which occurred as a result of change of mind of the agents of C with the linear orders R_{i_1}, \dots, R_{i_k} being replaced by them with linear orders $R'_{i_1}, \dots, R'_{i_k}$, if $F(\vec{R}')R_{i_s}F(\vec{R})$ for all $s = 1, 2, \dots, k$.

In other words, a profile is manipulable by a coalition C if agents of C can misrepresent their preferences in such a way that every member of C will benefit from the change.

A scoring rule F is characterised by the sequence of scores $v_1 \geq v_2 \geq \dots \geq v_m$. For each profile $\vec{R} \in \mathcal{L}(A)^n$ and for every alternative $a \in A$, we can define the score of a , denoted $Sc_F(\vec{R}, a)$, which can be computed as $Sc_F(\vec{R}, a) = \sum_{\ell=1}^m v_\ell i_\ell$, where the number i_k shows how many times the alternative a was ranked k th. The alternative with the highest score is declared the winner. If more than one alternative has the greatest score, an agreed-upon tie-breaking procedure is employed to determine the winner. For most models of the behaviour of the population, the probability of a tie tends to zero as $n \rightarrow \infty$. Note that it is possible, and convenient, to assume that $v_1 = 1$ and $v_m = 0$.

We can formalise this situation as follows. We write the linear order R_i as

$$a_{P_i(1)} > a_{P_i(2)} > \dots > a_{P_i(m)}$$

where $P_i \in \text{Sym}(m)$ is a permutation from the symmetric group $\text{Sym}(m)$ on $\{1, 2, \dots, m\}$. This way we may assume that each agent i chooses a permutation $P_i \in \text{Sym}(m)$. Then the vector total $X_n = \sum_{i=1}^n P_i v$ is found, whose j th component is the score $Sc_F(\vec{R}, a_j)$, and the winning candidate is the one corresponding to the greatest element of X_n .

Although the score vector v is not random in nature, nevertheless the following two characteristics of this vector will play a crucial role. We set

$$\bar{v} = \frac{1}{m} \sum_{j=1}^m v_j, \quad \text{and} \quad \sigma_v^2 = \frac{1}{m} \sum_{j=1}^m (v_j - \bar{v})^2.$$

The three most popular scoring rules are as follows.

Example 1. *Plurality voting:* $v = (1, 0, \dots, 0)$. We have $\bar{v} = 1/m$, and $\sigma_v^2 = (m-1)/m^2$.

Example 2. *Borda's rule:* $v = (1, \frac{m-2}{m-1}, \dots, \frac{1}{m-1}, 0)$. We have $\bar{v} = 1/2$, and $\sigma_v^2 = (m+1)/12(m-1)$.

Example 3. *Antiplurality voting:* $v = (1, \dots, 1, 0)$. We have $\bar{v} = m-1/m$, and $\sigma_v^2 = (m-1)/m^2$.

Example 4. *k -Approval voting:* $v = (1, \dots, 1, 0, \dots, 0)$ (k ones). We have $\bar{v} = k/m$, and $\sigma_v^2 = k(m-k)/m^2$.

Plurality and antiplurality rules are particular cases of the k -approval voting for $k = 1$ and $k = m-1$, respectively. These scoring rules will be further called *classical*.

The difference between k -Approval voting and Approval voting has to be emphasised. Under Approval voting any voter may approve as many alternatives as she wishes, while under k -Approval voting it is compulsory that any voter approves her exactly k best alternatives. Thus Approval voting is not a scoring rule. We refer the reader to the book by S. Brams and P. Fishburn [3] and to the more recent survey of its applications by S. Brams [4].

3 The Winning Margin

This section addresses the problem: if the voters choose permutations independently and uniformly at random from $\text{Sym}(m)$ (which corresponds to the IC conjecture), what is the asymptotic behaviour of the winning margin (i.e. the difference between the greatest and second-greatest elements of X_n)? The central limit theorem is used to find the asymptotics.

Proposition 1 (Central limit result.).

$$\frac{X_n - n\bar{v}1}{\sigma_v\sqrt{n}} \xrightarrow{D} N(0, \Sigma), \quad (1)$$

where Σ is the $m \times m$ matrix with diagonal elements equal to 1 and off-diagonal elements equal to $\frac{-1}{m-1}$.

Proof. The $P_i v$ are i.i.d. random vectors with mean $\bar{v}1$ and covariance matrix $\sigma_v^2 \Sigma$. The result follows by the classical central limit theorem (see e.g. [6], p. 170). \square

Define $\delta: \mathbb{R}^m \rightarrow [0, \infty)$ such that $\delta(x)$ is the difference between the greatest and second-greatest elements of x . Note that δ is continuous, with the properties $\delta(ax) = a\delta(x)$ for $a \geq 0$; $\delta(x + b1) = \delta(x)$ for $b \in \mathbb{R}$. We can thus apply δ to both sides of (1) to obtain

$$\frac{\delta(X_n)}{\sigma_v\sqrt{n}} \xrightarrow{D} \delta(Y), \quad (2)$$

where $Y \sim N(0, \Sigma)$.

Proposition 2 (Convergence of expectation). *We also have*

$$\frac{E[\delta(X_n)]}{\sigma_v\sqrt{n}} \rightarrow E[\delta(Y)]. \quad (3)$$

Proof. For random variables X'_n, Y' with the same distributions as X_n, Y (but perhaps defined on a different probability space) the convergence in (2) takes place in the almost-sure sense. We will also have $\delta(X'_n)/\sigma_v\sqrt{n} \rightarrow \delta(Y')$ in the L_1 sense (from which the desired result will follow) provided we can show that the family $\{\delta(X_n)/\sqrt{n}\}_{n=1}^\infty$ is uniformly integrable. For this, it is enough to show that $\{\|X_n - n\bar{v}1\|_2/\sqrt{n}\}_{n=1}^\infty$ is uniformly integrable, since $\delta(x) = \delta(x - b1) \leq 2\|x - b1\|_\infty \leq 2\|x - b1\|_2$ for all $x \in \mathbb{R}^m$ and $b \in \mathbb{R}$. But this holds since

$$E \left[\frac{\|X_n - n\bar{v}1\|_2^2}{n} \right] = \frac{1}{n} \sum_{j=1}^m E \left[\left(\sum_{i=1}^n (P_i v)_j - n\bar{v} \right)^2 \right] = m\sigma_v^2,$$

i.e. the family $\{\|X_n - n\bar{v}1\|_2/\sqrt{n}\}$ is bounded in mean square. \square

Proposition 3 (The distribution of $\delta(Y)$). *The random vector Y has the same multivariate normal distribution as $\sqrt{\frac{m}{m-1}}(Z - \bar{Z}1)$, where $Z \sim N(0, I_m)$ and $\bar{Z} = \frac{1}{m} \sum_{j=1}^m Z_j$. In particular,*

$$E[\delta(Y)] = \sqrt{\frac{m}{m-1}} E[\delta(Z)].$$

Proof. The distribution of Y is $N(0, \Sigma)$. To see that $\sqrt{\frac{m}{m-1}}(Z - \bar{Z}1)$ also has this distribution, one only has to verify that $\text{Cov}(Z_j - \bar{Z}, Z_k - \bar{Z}) = \delta_{jk} - 1/m$ and $E[Z_j - \bar{Z}] = 0$. \square

The values of $E_m = E[\delta(Y)]$ for various m can be found from several published tables of the expected order statistics of normal variables; e.g. those in [10]. This allows the construction of the following table:

m	3	4	5	6	7	8	9	10	16
E_m	1.036	0.845	0.747	0.685	0.643	0.611	0.586	0.567	0.4971

Theorem 1. *The average winning margin $WM(n, m, v) = E[\delta(X_n)]$ is asymptotic to $E_m\sigma_v\sqrt{n}$ as $n \rightarrow \infty$.*

Proof. Follows from Propositions 1–3. \square

4 Threshold Coalition Size

Given a profile, in order to determine the minimum manipulating coalition size for this profile we have to know the winning margin and how the scores may change during the manipulation attempt. Suppose we are given the

scoring rule with a score vector $v = (v_1, \dots, v_m)$. Suppose also that at a certain profile the alternative a wins, while b has the second-best score. A voter wants to manipulate in favour of b . Then b must have a higher rank on the ballot of this voter. The best that the voter can do is to move b to the first position on her ballot and move a to the last position. This will reduce the winning margin by $(v_1 - v_i) + (v_j - v_m)$, where i and j are the original positions of b and a respectively; note $i \leq j$. The greatest amount by which any single voter will be able (and willing) to reduce the winning margin is thus

$$d_v = \max_i [(v_1 - v_i) + (v_{i+1} - v_m)].$$

For the Borda rule, $d_v = (m - 2)/(m - 1)$. For plurality and k -approval voting, $d_v = 1$.

Let the *threshold coalition size* M_v be (the random variable) given by

$$M_v = \frac{\delta(X_n)}{d_v},$$

where, as in the last section, $\delta(X_n)$ is the winning margin. Any coalition capable of manipulating the outcome of the election must have at least $\lceil M_v \rceil$ members.

A note of caution must be given here. It is not true that any coalition of more than $\lceil M_v \rceil$ members can manipulate at the given profile. This coalition is capable of reversing the gap between the best and the second best alternative but it is possible that the third (and highly undesirable) alternative will overtake both and ruin the manipulation attempt. This can be illustrated on the following example:

Example 5. *Let us consider the Borda rule B and the following profile:*

R_1	R_2	R_3	R_4	R_5	R_6
a	a	b	b	c	c
b	b	a	c	a	b
c	c	c	a	b	a

As $Sc_B(\vec{R}, a) = 6$, $Sc_B(\vec{R}, b) = 7$ and $Sc_B(\vec{R}, c) = 5$, the winner is b . The coalition $\{1, 2\}$ may attempt to manipulate to get a chosen submitting linear orders R'_1 and R'_2 , where they swap the order of b and c . By doing so, they will reverse the winning margin between b and a but their attempt will fail since c will be chosen, which is their worst alternative.

Nevertheless M_v is a very important parameter of the rule. We will show that for classical scoring rules this number asymptotically coincides with the minimum size of the coalition that is capable of manipulation.

Theorem 2. *The average threshold coalition size, $ATC(n, m, v) = E[M_v]$, is asymptotic to $C(m, v)\sqrt{n}$ as $n \rightarrow \infty$, where $C(m, v) = E_m\sigma_v/d_v$.*

Proof. Follows from Theorem 1 and the arguments given at the beginning of this section. \square

5 Minimum Manipulating Coalition Size

Our next goal is to prove that in a large electorate it is almost certain that a coalition will exist that can reverse the margin between the best and second-best alternatives. Moreover, for the classical voting rules it will be of size $\lceil M_v \rceil$ or $\lceil M_v \rceil + 1$ (depending on the tie-breaking procedure). First, we will show that, for any constant t and for any linear order $R \in \mathcal{L}(A)$, it is almost certain that, for n large enough, at least $t\lceil M_v \rceil$ voters will submit this linear order. To see this, note that the number of voters who submit R will be of order n , but $\lceil M_v \rceil$ is of order only \sqrt{n} . The following result states this fact more formally.

Theorem 3. *Let $R \in \mathcal{L}(A)$ be a linear order and $\vec{R} = (R_1, \dots, R_n)$ be a random profile. Let $t > 0$ be a positive number. Then, for any scoring rule with score vector v , there is a number $\beta > 0$ such that*

$$P(\#\{i \mid R_i = R\} \leq t\lceil M_v \rceil) \geq 1 - 2e^{-\beta n}$$

for all large enough n . That is, the probability that a coalition of like-minded voters, who submit linear order R , of size $t\lceil M_v \rceil$ exists converges to 1 with exponential rate as $n \rightarrow \infty$.

Proof. Let D_i be the random variable which is 1 if voter i submits R and 0 otherwise. Let $M' = \sum_{i=1}^n D_i$ be the number of voters who submit R ; the claim is that $M' \geq tM_v$ with high probability. Let $p = E[D_i]$. The IC conjecture states that $p = 1/m! > 0$. Choose $\alpha \in (0, p)$; then

$$\{M' < \alpha n\} \subseteq \left\{ \frac{1}{n} \sum_{i=1}^n (D_i - p) \leq \alpha - p \right\}$$

Also,

$$\{M_v > \alpha n/t\} = \left\{ \delta \left(\frac{X_n}{n} - \bar{v}1 \right) > \alpha d_v/t \right\} \subseteq \left\{ \frac{X_n}{n} - \bar{v}1 \in \delta^{-1}([\alpha d_v/t, \infty)) \right\}.$$

Since $(-\infty, \alpha - p]$ and $K = \delta^{-1}([\alpha d_v/t, \infty))$ are closed sets which do not include the origin, Cramér's large-deviation theorem ([7], p.36) gives that for some $\beta > 0$,

$$P\left(\frac{1}{n}\sum_{i=1}^n(D_i - p) \leq \alpha - p\right) < e^{-\beta n} \quad \text{and} \quad P\left(\frac{X_n}{n} - \bar{v}1 \in K\right) < e^{-\beta n}$$

for all large enough n . It follows that

$$P(M' \geq tM_v) \geq P(M' \geq \alpha n \text{ and } M_v \leq \alpha n/t) \geq 1 - 2e^{-\beta n}$$

for all large enough n . □

The following theorem was effectively proved in [12]. Here we prove a slightly stronger statement.

Theorem 4 (Kim-Roush, 1996). *Any scoring voting rule F , with the exception of antiplurality, is coalitionally manipulable with limiting probability 1, i.e. the probability that a random profile is coalitionally manipulable for F converges to 1 with exponential rate as $n \rightarrow \infty$. If F is one of the classical scoring rules with the score vector v , then the probability of existence of manipulating coalition of size $M_v + 1$ also converges to 1 with exponential rate.*

Proof. Let $\vec{R} = (R_1, \dots, R_n)$ be a random profile. Let $F(\vec{R}) = w$ and s is the alternative with the second-best score. Let v be the score vector of F and let $d = v_1 - v_{n-1} = 1 - v_{n-1}$. Since F is different from the antiplurality, $v_{n-1} \neq 1$ and $d > 0$.

Let $t > d_v/d$. By Theorem 3, among R_i 's there will be at least $t[M_v]$ of linear orders R for which w occupies the last place and s occupies second to last. Moving s to the top place without changing the order of other alternatives will add $dt[M_v]$ points to the score of s while the score of w will remain the same and the scores of other alternatives will not increase. As $dt[M_v] > \delta(X_n)$, the new winner will be s and this manipulation attempt will be successful.

If F is classical and different from antiplurality, then $d = d_v$ and as we seen above \vec{R} will be almost surely manipulable by a coalition of size $[M_v] + 1$. □

The antiplurality A is the exception indeed. In the three-alternative case the limiting probability of coalitional manipulability will be $1/2$. If alternatives are denoted a, b, c , then a profile with $Sc_A(a) > Sc_A(b) > Sc_A(c)$ is manipulable if $Sc_A(a) + Sc_A(c) < 2Sc_A(b)$ and not manipulable

if $Sc_A(a) + Sc_A(c) > 2Sc_A(b)$. Due to the symmetry these two events have equal probabilities. Kim and Roush [12] gave the formula for calculating the limiting probability for $m > 3$.

Let F be a scoring rule with a score vector v . Let the *minimum manipulating coalition size* MC_v be (the random variable) equal to the minimum size of a coalition capable of manipulation, if the random profile is manipulable, and ∞ if it is not. Let average minimum manipulating coalition size $AMMC(n, m, v)$ be the expectancy of MC_v conditional on the event that the random profile is manipulable.

We recap the results of this section in the following

Theorem 5. *For any scoring rule different from the antiplurality,*

$$AMMC(n, m, v) \geq ATC(n, m, v).$$

The average minimum manipulating coalition size $AMMC(n, m, v)$ and average threshold coalition size $ATC(n, m, v)$, are asymptotically equal and, as $n \rightarrow \infty$, both are asymptotic to $C(m, v)\sqrt{n}$, where $C(m, v) = E_m\sigma_v/d_v$.

Proof. By Theorem 4 we may discard nonmanipulable profiles. Hence, the first part is clear since $MC_v \geq M_v$. For all rules but antiplurality the remainder of the theorem follows from Theorems 2 and 4. \square

6 A Drawback of the Average Minimum Manipulating Coalition Size

We will illustrate this drawback in the case of $m = 3$ alternatives, but apparently it appears for all $m \geq 3$.

Consider the scoring rule $v = (1, 1 - p, 0)$, where $0 \leq p \leq 1$. For a given profile, denote the winning candidate a , the second-placegetter b , and the third-placegetter c . Let $d_1 = \delta(X_n)$ be the winning margin of a over b , and $d_2 = \delta(-X_n)$ be the “second margin” between b and c . Two types of manipulation in favour of b are possible:

- A voter who prefers c to b to a may mis-represent his preference as b, c, a . The effect of this will be to reduce the winning margin by p , while increasing the second margin by $2p$.
- A voter who prefers b to a to c may mis-represent his preference as b, c, a . The effect of this will be to reduce both the winning and second margins by $1 - p$.

Suppose a coalition of voters contains x_1 voters of the first type and x_2 voters of the second type. Successful manipulation by this coalition requires

$$\begin{aligned} d_1 - px_1 - (1-p)x_2 &\leq 0, \\ d_2 + 2px_1 - (1-p)x_2 &\geq 0. \end{aligned}$$

If $0 < p \leq 1/2$, then the smallest coalition that can successfully manipulate has

$$x_1 = 0, \quad x_2 = \frac{d_1}{1-p}$$

when $d_1 \leq d_2$, and

$$x_1 = \frac{d_1 - d_2}{3p}, \quad x_2 = \frac{2d_1 + d_2}{3(1-p)}$$

when $d_1 \geq d_2$. If, on the other hand, $1/2 \leq p \leq 1$, then the minimal coalition has

$$x_1 = \frac{d_1}{p}, \quad x_2 = 0.$$

Let us denote $x_+ = x$, if $x \geq 0$ and $x_+ = 0$, otherwise. Then the minimum manipulating coalition size can be written as

$$MC_v = \frac{1}{\max(p, 1-p)} d_1 + \frac{1}{3} \left(\frac{1}{p} - \frac{1}{1-p} \right)_+ (d_1 - d_2)_+$$

in all cases where $0 < p \leq 1$. (There is a small probability that insufficiently many like-minded voters will exist to form such a minimal coalition. But by Theorem 3, this possibility can be neglected for our purposes.)

If we define $\rho : \mathbb{R}^3 \rightarrow [0, \infty)$ by $\rho(x) = (\delta(x) - \delta(-x))_+$, the excess of the winning margin over the second margin, and note that for these rules $d_v = \max(p, 1-p)$, then we can write

$$MC_v = \frac{\delta(X_n)}{d_v} + \frac{1}{3} \left(\frac{1}{p} - \frac{1}{1-p} \right)_+ \rho(X_n).$$

Note that the first term of this expression is the threshold coalition size M_v .

The average minimum manipulating coalition size can be taken to be the expectation of this quantity. (Strictly speaking, this expectation should be conditional on manipulability of the profile. But since the probability of this event converges rapidly to 1 – by Theorem 3 again – the conditional expectation can be replaced by an unconditional one for our purposes.) Also, by the same arguments as used in section 3, we have

$$\frac{E[\delta(X_n)]}{\sigma_v \sqrt{n}} \rightarrow \sqrt{\frac{m}{m-1}} E[\delta(Z)] \quad \text{and} \quad \frac{E[\rho(X_n)]}{\sigma_v \sqrt{n}} \rightarrow \sqrt{\frac{m}{m-1}} E[\rho(Z)],$$

where $Z \sim N(0, I_m)$. The average minimum manipulating coalition size thus satisfies

$$\frac{AMMC(n, 3, v)}{\sigma_v \sqrt{n}} \approx \frac{E[MC_v]}{\sigma_v \sqrt{n}} \rightarrow \frac{E_3}{d_v} + \frac{1}{\sqrt{6}} \left(\frac{1}{p} - \frac{1}{1-p} \right)_+ E[\rho(Z)].$$

When $1/2 \leq p \leq 1$ the second term vanishes. For such rules, the average threshold coalition size and the average minimum manipulating coalition size have the same asymptotics. But when $0 < p < 1/2$, the value of $\gamma(v) = \lim_{n \rightarrow \infty} AMMC(n, 3, v)/\sqrt{n}$ is greater than that of $\lim_{n \rightarrow \infty} ATC(n, 3, v)/\sqrt{n}$. For small p , it is much greater, and indeed $\lim_{p \rightarrow 0} \gamma(v) = \infty$.

Now let us consider the case $p = 0$: the anti-plurality rule. For this rule, only the voters of the first type have power to manipulate. If $d_1 < d_2$, no manipulation is possible; otherwise the smallest manipulating coalition has

$$x_1 = 0, \quad x_2 = d_1.$$

Hence

$$\begin{aligned} AMMC(n, 3, v) &= E[MC_v | d_1 \geq d_2] \\ &= E[d_1 | d_1 \geq d_2] \\ &= E[d_1 1_{d_1 \geq d_2}] / P(d_1 \geq d_2) \\ &= 2E[\delta(X_n) 1_{\delta(X_n) \geq \delta(-X_n)}], \end{aligned}$$

and so

$$\gamma(v) = \lim_n \frac{AMMC(n, 3, v)}{\sqrt{n}} = 2\sigma_v \sqrt{3/2} E[\delta(Z) 1_{\delta(Z) \geq \delta(-Z)}],$$

which is a finite number.

If we adopt the average minimum manipulating coalition size as a measure of goodness of the rule, then we will have to accept that, the closer the rule to antiplurality the better it becomes, and that there are no optimal rules. This would be clearly an absurd conclusion. Hence, no matter how attractive it may seem, the average minimum manipulating coalition size is not an appropriate measure for comparing rules. We suggest that the average threshold coalition size should be used instead.

7 Optimal Scoring Rules

In this section we will determine the rules optimal in relation to asymptotic average threshold coalition size. For $m = 3, 4$ the optimal rule is Borda Count and for $m \geq 5$ the optimal rule is $\lfloor m/2 \rfloor$ -approval voting.

By Theorem 5, we may assume that the quality of the rule is determined by the ratio σ_v/d_v , or on the number $C(m, v) = E_m \sigma_v/d_v$ which we will analyse in this section.

Theorem 6. *The asymptotic average minimum manipulating coalition size is maximised among all scoring rules by the Borda rule when the number m of alternatives is 3 or 4, and by $\lfloor m/2 \rfloor$ -approval voting when $m \geq 5$.*

Proof. We must show that the claimed rules maximise the ratio σ_v/d_v (or equivalently, σ_v^2/d_v^2) among all score vectors v .

Note that

$$\sigma_v^2 = \frac{1}{m} \sum_{i=1}^m (v_i - \bar{v})^2 = \frac{1}{2} E [(U - V)^2] = \frac{1}{m^2} \sum_{i < j} (v_i - v_j)^2,$$

where U and V are independent copies of $(P_1 v)_1$. We recall that, $v_1 = 1$ and $v_m = 0$. If we parametrize the score vectors by setting $v_i = \sum_{j=i}^{m-1} t_j$, where $t_1, \dots, t_{m-1} \geq 0$ and $\sum_{j=1}^{m-1} t_j = 1$, then

$$\sigma_v^2 = S(t) = \frac{1}{m^2} \sum_{i \leq j} \left(\sum_{k=i}^j t_k \right)^2.$$

Let us first consider all the score vectors v with a fixed value of d_v . Among these, we only have to maximise σ_v . The set

$$V_\alpha = \left\{ (t_1, \dots, t_{m-1}) : \text{all } t_i \geq \alpha \text{ and } \sum_{i=1}^{m-1} t_i = 1 \right\}$$

includes a representation of all the score vectors with $d_v = 1 - \alpha$. Since S is a convex function and V_α is a convex set, the maximum value of S on V_α must be attained at an extremal point of V_α , i.e. at one of the points $t^{(i)}(\alpha)$ with every component except the i th equal to α (and the i th component equal to $1 - (m - 2)\alpha$). In particular, the maximum value of σ_v among all score vectors with $d_v = 1 - \alpha$ is attained by a v corresponding to one of the $t^{(i)}(\alpha)$.

Now let us see that the maximum is achieved at $t^*(\alpha) = t^{(\lfloor m/2 \rfloor)}(\alpha)$, and also at $t^{(\lceil m/2 \rceil)}(\alpha)$. Observe that $S(t^{(i)}(\alpha)) = S(t^{(m-i)}(\alpha))$ by symmetry, so it is enough to note that when $i < \lfloor m/2 \rfloor$,

$$S(t^{(i)}(\alpha)) - S(t^{(i+1)}(\alpha)) < 0. \quad (4)$$

Since $\alpha \leq \frac{1}{m-1}$, we have $1 - (m - 2)\alpha \geq \alpha$. Now a direct calculation shows that

$$S(t^{(i)}(\alpha)) - S(t^{(i+1)}(\alpha)) = \frac{1}{m^2} \left\{ \sum_{r=1}^i ((1 - (m - 2)\alpha + (r - 1)\alpha)^2 - (r\alpha)^2) \right.$$

$$\begin{aligned}
& + \left. \sum_{r=1}^{m-i-1} \left((r\alpha)^2 - (1 - (m-2)\alpha + (r-1)\alpha)^2 \right) \right\} \\
& = \sum_{r=i+1}^{m-i-1} \left((r\alpha)^2 - (1 - (m-2)\alpha + (r-1)\alpha)^2 \right) < 0.
\end{aligned}$$

It now remains to find

$$\max \left\{ \frac{S(t^*(\alpha))}{(1-\alpha)^2} : 0 \leq \alpha \leq \frac{1}{m-1} \right\}. \quad (5)$$

For $m = 3$, we have

$$\frac{S(t^*(\alpha))}{(1-\alpha)^2} = \frac{S(1-\alpha, \alpha)}{(1-\alpha)^2} = \frac{2(1-\alpha+\alpha^2)}{9(1-\alpha)^2} = \frac{2}{9} \left(1 + \frac{\alpha}{(1-\alpha)^2} \right).$$

Since this is an increasing function of α , its maximum is achieved when α takes its greatest possible value of $1/2$. This corresponds to the Borda rule.

For $m = 4$, we have

$$\frac{S(t^*(\alpha))}{(1-\alpha)^2} = \frac{S(\alpha, 1-2\alpha, \alpha)}{(1-\alpha)^2} = \frac{1-2\alpha+2\alpha^2}{4(1-\alpha)^2} = \frac{1}{4} \left(1 + \frac{\alpha^2}{(1-\alpha)^2} \right).$$

Since this is again an increasing function of α , its maximum is achieved when α takes its greatest possible value of $1/3$. This also corresponds to the Borda rule.

Now we turn our attention to the case $m \geq 5$. Note that in general

$$\begin{aligned}
S(t^*(\alpha)) &= \frac{1}{m^2} \left\{ \sum_{i < j < \lfloor m/2 \rfloor} ((j-i+1)\alpha)^2 + \sum_{\lfloor m/2 \rfloor < i \leq j} ((j-i+1)\alpha)^2 \right. \\
&\quad \left. + \sum_{i \leq \lfloor m/2 \rfloor \leq j} (1 - (m-2-j+i)\alpha)^2 \right\} \\
&= c_m \left(1 - (m-2)\alpha + \left(\frac{2+e_m-3m+m^2}{3} \right) \alpha^2 \right),
\end{aligned}$$

where $c_m = (1 - e_m/m^2)/4$ and e_m is 1 if m is odd, 0 if m is even. To complete our proof, we must show that the maximum in (5) is attained when $\alpha = 0$, i.e. that $S(t^*(\alpha))/(1-\alpha)^2 \leq S(t^*(0)) = c_m$ for $0 \leq \alpha \leq 1/(m-1)$. To this end, note that

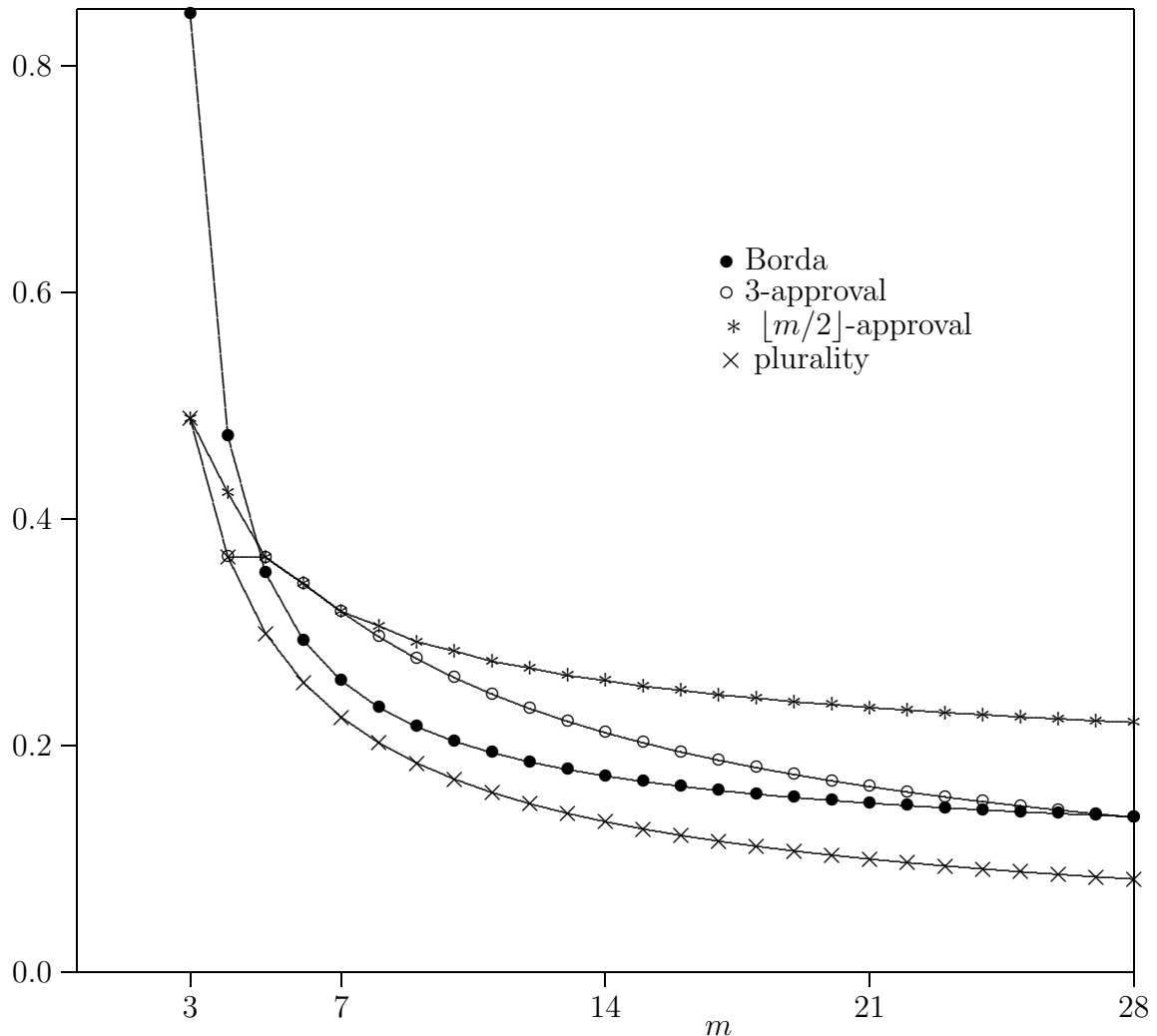
$$S(t^*(\alpha)) - c_m(1-\alpha)^2 = c_m \alpha \left(-(m-4) + \left(\frac{-1+e_m-3m+m^2}{3} \right) \alpha \right).$$

Since $\frac{1}{3}(-1 + e_m - 3m + m^2) \geq 0$ for $m \geq 5$, for these values of m we obtain

$$\begin{aligned} S(t^*(\alpha)) - c_m(1 - \alpha)^2 &\leq c_m \alpha \left(-(m - 4) + \left(\frac{-1 + e_m - 3m + m^2}{3} \right) \frac{1}{m - 1} \right) \\ &= \frac{c_m \alpha}{3(m - 1)} (e_m - 13 + 12m - 2m^2) \leq 0. \end{aligned}$$

This ends the proof. \square

The following graph and table display the numbers $C(m, v)$ for small m for the two most popular scoring rules plurality and Borda, and also for $\lfloor m/2 \rfloor$ -approval and 3-approval voting.



The following table gives some numerical values of the numbers $C(m, v)$ for small m :

	3	4	5	6	7	8	9	10
Borda	0.846	0.472	0.352	0.292	0.257	0.233	0.216	0.203
Plurality	0.488	0.366	0.299	0.255	0.225	0.202	0.184	0.170
3-Appr.		0.366	0.366	0.343	0.318	0.296	0.276	0.260
$\lfloor m/2 \rfloor$ -Appr.	0.488	0.423	0.366	0.343	0.318	0.306	0.291	0.284

We included 3-Approval voting in this table for the following reason. In the recent presidential elections in France, in which there were 16 candidates, the following experiment was conducted [13]. The experiment was run in six places: in a village, called Gy-les-Nonnains and in five voting posts in Orsay, a suburb of Paris. In these places voters had a possibility, after casting their official ballot, to vote unofficially according to the approval voting rule. In both places the average number of approvals on a ballot was three. So 3-Approval voting may be the best approximation to Approval voting among the scoring rules. Comparing it with Borda, we see that Borda eventually gets better than 3-Approval but it happens only for $m \geq 28$.

8 Conclusion

We see that there are no optimal rules relative to the asymptotic average minimum manipulating coalition size. This characteristic is biased towards the antiplurality and should be used with great caution. The average threshold coalition size is maximised among all scoring rules by the Borda rule not only for the three-alternative case but also for the case of four alternatives. At the same time, when the number of alternatives is larger than four the $\lfloor m/2 \rfloor$ -approval voting is preferable. It is interesting that in the Saari's geometric model for $m = 4$ the 2-approval voting is better than Borda in respect to micro manipulation [18] and Saari expected Borda to fare poorly relative to coordinated macro manipulation. Although the difference in models should be taken into account, our results do not seem to confirm these expectations.

It is hard to imagine $\lfloor m/2 \rfloor$ -approval voting implemented in practice. The best approximation to this rule would be Majoritarian Compromise which for large societies behaves itself as $\lfloor m/2 \rfloor$ -approval voting [20].

Some experimental evidence [13] suggests that 3-Approval voting might be the best approximation among scoring rules to the classical Approval voting. If so, the characteristics of Approval voting seem to be quite good since Borda surpasses 3-Approval only for $m \geq 28$.

It is interesting to compare these results with the results obtained by Chamberlin in [5] using computer simulation. The four rules plurality, Borda, Hare and Coombs were tested using the IC conjecture in three-alternative

elections and for those profiles for which all four rules were manipulable (which is about 30% of all profiles), the average minimum manipulating coalition size was recorded in the following table.

Voting system	21 voters	1000 voters
Plurality	2.4	12.4
Borda	2.3	15.4
Hare	1.5	5.9
Coombs	2.5	11.6

We see that for 21 voters the result of comparison is inconclusive while for 1000 voters Borda is clearly the best.

Since for classical rules the asymptotic average minimum coalition size is the same as the asymptotic average threshold coalition size, we can compare Chamberlin's results with ours. For 1000 voters our results give us 15.4 for the plurality and 26.8 for the Borda rule. The difference is especially significant for the Borda rule. One possible explanation to this discrepancy is that in Chamberlin's simulations only profiles manipulable for all four rules, including Hare's rule, were considered. Since Hare's rule is manipulable only in 30% of all cases, this can be a rather restricted set of profiles which are more prone to manipulation than the majority of profiles.

References

- [1] Aleskerov F and Kurbanov E (1999) Degree of manipulability of social choice procedures. In: *Current Trends in Economics*. Eds: Alkan, A., Aliprantis, C.D. and Yannelis, N.C. Springer, 13–27
- [2] Berg S and Lepelley D (1994) On Probability models in voting theory. *Statistica Neerlandica* 48, 133–146
- [3] Brams S and Fishburn P (1982) *Approval voting*. Birkhauser
- [4] Brams S (2002) From theory to practice: the mixed success of Approval Voting, mimeo, New York University.
- [5] Chamberlin J (1985) Investigation into the relative manipulability of four voting systems. *Behavior Science* 30: 195–203
- [6] Durrett R (1996) *Probability: theory and examples*, 2nd ed. Duxbury Press
- [7] Dembo A and Zeitouni O (1993) *Large deviations techniques*. Jones and Bartlett
- [8] Favardin P Lepelley D and Serias J (2002) Borda rule, Copeland method and strategic manipulation. *Review of Economic Design* 7, 213 –228
- [9] Gibbard A (1973) Manipulation of voting schemes: A general result. *Econometrica* 41:587–601
- [10] Harter LH (1961) Expected values of normal order statistics. *Biometrika* 48: 151–165
- [11] Kelly J (1993) Almost all social choice rules are highly manipulable, but few aren't. *Social Choice and Welfare* 10:161–175
- [12] Kim KH and Roush FW (1996) Statistical manipulability of social choice functions. *Group Decision and Negotiation* 5:263–282
- [13] Laslier J-P (2003) Spatial approval voting, to appear.
- [14] Lepelley D and Mbih B (1987) The proportion of coalitionally unstable situations under the plurality rule. *Economic Letters* 24, 311–315
- [15] Lepelley D and Mbih B (1994) The vulnerability of four social choice functions to coalitional manipulation of preferences. *Social Choice and Welfare*, 11: 253–265

- [16] Lepelley D and Valognes F (2002) Voting rules, manipulability and social homogeneity. *Public Choice*, forthcoming
- [17] Nitzan S (1985) The vulnerability of point-voting schemes to preference variation and strategic manipulation. *Public Choice* 47: 349–370
- [18] Saari D (1990) Susceptibility to manipulation. *Public Choice* 64:21–41
- [19] Satterthwaite MA (1975) Strategy-proofness and Arrow's conditions: Existence and correspondence theorems for voting procedures and social welfare functions. *Journal of Economic Theory* 10:187–217
- [20] Slinko A (2002) The Majoritarian Compromise in Large Societies, *Review of Economics Design* 7(3): 343–350.
- [21] Smith DA (1999) Manipulability measures of common social choice functions. *Social Choice and Welfare* 16: 639–661